## MASSACHUSETTS INSTITUTE OF TECHNOLOGY

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## Exercise 1.

- (a) If  $X_1 \sim Cauchy(0, \gamma_1)$ ,  $X_2 \sim Cauchy(0, \gamma_2)$ , and they are independent, then  $X_1 + X_2 \sim Cauchy(0, \gamma_1 + \gamma_2)$ .
- (b) If  $X \sim Cauchy(0, \gamma)$ , then  $\alpha X \sim Cauchy(0, \alpha \gamma)$ , for all  $\alpha > 0$ .
- (c) Let  $\{X_n\}$  be a sequence of i.i.d. random variables, with  $X_1 \sim Cauchy(0, \gamma)$ . Then,

$$\frac{X_1 + \dots + X_n}{n} \sim Cauchy(0, \gamma),$$

for all n.

## **Solution:**

(a) We have

$$\phi_{X_1+X_2}(t) = \phi_{X_1}(t)\phi_{X_2}(t) = \exp\left(-\gamma_1|t|\right)\exp\left(-\gamma_2|t|\right) = \exp\left(-(\gamma_1+\gamma_2)|t|\right),$$

which corresponds to a  $Cauchy(0, \gamma_1 + \gamma_2)$ .

(b) We have

$$\phi_{\alpha X}(t) = \phi_X(\alpha t) = \exp(-\alpha \gamma |t|),$$

which corresponds to a  $Cauchy(0, \alpha \gamma)$ .

(c) We have

$$\phi_{\frac{X_1 + \dots + X_n}{n}}(t) = \prod_{k=1}^n \phi_{X_k} \left(\frac{t}{n}\right)^n$$
$$= \prod_{k=1}^n \exp\left(-\gamma \left|\frac{t}{n}\right|\right)$$

which corresponds to a  $Cauchy(0, \gamma)$  for all n.

**Exercise 2.** Let  $\{X_n\}$  be a sequence of random variables, such that  $\mathbb{E}[X_n] = 0$  and  $Var(X_n) \leq \sigma^2$  for all n, and such that  $Cov(X_i, X_j) \to 0$  when  $|i - j| \to \infty$ . Then,

$$S_n = \frac{X_1 + \dots + X_n}{n} \xrightarrow{i.p.} 0.$$

**Solution:** For any  $\epsilon > 0$ , Chebyshev's inequality implies that

$$\mathbb{P}(|S_n| \ge \epsilon) \le \frac{Var(S_n)}{\epsilon^2} = \frac{1}{n^2 \epsilon^2} \sum_{i=1}^n \sum_{j=1}^n Cov(X_i, X_j).$$

Since  $Cov(X_i,X_j) \to 0$  when  $|i-j| \to \infty$ , then for every  $\delta > 0$ , there exists  $N_\delta$  such that  $|Cov(X_i,X_j)| \le \delta$  for all i,j such that  $|i-j| > N_\delta$ . Thus, we have

$$\begin{split} \frac{1}{n^2\epsilon^2} \sum_{i=1}^n \sum_{j=1}^n Cov(X_i, X_j) &= \frac{1}{n^2\epsilon^2} \sum_{i=1}^n \left( \sum_{j: |i-j| \leq N_\delta} Cov(X_i, X_j) + \sum_{j: |i-j| > N_\delta} Cov(X_i, X_j) \right) \\ &\leq \frac{1}{n^2\epsilon^2} \sum_{i=1}^n \left( \sum_{j: |i-j| \leq N_\delta} \sigma^2 + \sum_{j: |i-j| > N_\delta} \delta \right) \\ &\leq \frac{1}{n^2\epsilon^2} \Big[ n(2N_\delta + 1)\sigma^2 + n^2 \delta \Big] \\ &\leq \frac{2N_\delta + 1}{n\epsilon^2} + \frac{\delta}{\epsilon^2}. \end{split}$$

Taking limit as  $n \to \infty$ , we have

$$\lim_{n \to \infty} \mathbb{P}\left(|S_n| \ge \epsilon\right) \le \frac{\delta}{\epsilon^2}.$$

Finally, since this is true for all  $\delta > 0$ , we get

$$\lim_{n \to \infty} \mathbb{P}\left(|S_n| \ge \epsilon\right) = 0.$$

**Exercise 3.** Let  $\{X_n\}$  be a sequence of i.i.d. random variables such that  $X_1 \sim \mathcal{N}(0,1)$ . Let us define  $Y_k = X_1 + \cdots + X_k$ . Show that

$$\frac{Y_1 + \dots + Y_n}{n^{3/2}} \stackrel{d}{\longrightarrow} \mathcal{N}(0, 1/3).$$

**Solution:** Let us define

$$S_n = \sum_{k=1}^n Y_k.$$

Note that

$$S_n = \sum_{k=1}^{n} (n - k + 1) X_k.$$

Then, we have

$$\phi_{S_n}(t) = \prod_{k=1}^n \phi_{X_k}(tk)$$

$$= \prod_{k=1}^n \exp\left(-\frac{(tk)^2}{2}\right)$$

$$= \exp\left(-\frac{t^2}{2}\sum_{k=1}^n k^2\right)$$

$$= \exp\left[-\frac{t^2}{2}\left(\frac{n^3}{3} + \frac{n^2}{2} + \frac{n}{6}\right)\right],$$

and thus

$$\phi_{\frac{S_n}{n^{3/2}}}(t) = \phi_{S_n} \left( \frac{t}{n^{3/2}} \right)$$

$$= \exp \left[ -\frac{t^2}{2} \left( \frac{1}{3} + \frac{1}{2n} + \frac{1}{6n^2} \right) \right].$$

Finally, if  $S = \lim_{n \to \infty} \frac{S_n}{n^{3/2}}$ , we have

$$\begin{split} \phi_S(t) &= \lim_{n \to \infty} \phi_{\frac{S_n}{n^{3/2}}}(t) \\ &= \exp\left(-\frac{t^2}{6}\right), \end{split}$$

which corresponds to a  $\mathcal{N}(0, 1/3)$ .

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